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REASSESS V2.0: SOFTWARE FOR SINGLE- AND MULTI-SITE PROBABILISTIC SEISMIC HAZARD ANALYSIS.

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Abstract

Probabilistic seismic hazard analysis (PSHA) is generally recognized as the rational method to quantify the seismic threat. Classical formulation of PSHA goes back to the second half of the twentieth century, but its implementation can still be demanding for engineers dealing with practical applications. Moreover, in the last years, a number of developments of PSHA have been introduced; e.g., vector-valued and advanced ground motion intensity measure (IM) hazard, the inclusion of the effect of aftershocks in single-site hazard assessment, and multi-site analysis requiring the characterization of random fields of cross-correlated IMs. Although software to carry out PSHA has been available since quite some time, generally, it does not feature a user-friendly interface and does not embed most of the recent methodologies relevant from the earthquake engineering perspective. These are the main motivations behind the development of the practice-oriented software presented herein, namely REgionAl, Single-SitE and Scenario-based Seismic hazard analysis (REASSESS V2.0). In the paper, the seismic hazard assessments REASSESS enables are discussed, along with the implemented algorithms and the models/databases embedded at this stage of the software. Illustrative applications exploit the potential of the tool, which is available at http://wpage.unina.it/iuniervo/doc_en/REASSESS.htm.

25 1. Introduction

26 The classical (single-site) formulation of probabilistic seismic hazard analysis (PSHA) aims at evaluating the
27 rate of earthquakes causing exceedance of any arbitrary ground-motion intensity measure (IM) threshold (*im*)
28 at a site of interest (Cornell, 1968). PSHA lies at the basis of seismic risk assessment according to the
29 performance-based earthquake engineering paradigm (Cornell and Krawinkler, 2000) and serves for the
30 determination of seismic actions for structural design in several countries.

31 The probabilistic assessment of the seismic threat at a site is, in principle, not straightforward for structural
32 engineers because it requires the employment of models and skills they do not typically have at hand. For this
33 reason, during the last four decades, computer software to carry out PSHA have become available, starting
34 from EQRISK (McGuire, 1976). Other relevant codes are, for example, SEISRISK III (Bender and Perkins,
35 1987), OpenSHA (Field et al., 2003) and CRISIS (Ordaz et al., 2013); see Danciu et al. (2010). Recently, the
36 global earthquake model (GEM) foundation developed OpenQuake (Pagani et al., 2014) that has been adopted,
37 among others, within the EMME (Giardini et al., 2018) and SHARE (Giardini et al., 2013) hazard assessment
38 projects.

39 PSHA has been significantly extended since its introduction in the late sixties. For example, its classical
40 version refers to a scalar IM, while advanced structural assessment procedures may require hazard in terms of
41 vector-valued IMs (Baker and Cornell, 2006b) or, equivalently, development of *conditional hazard* for
42 secondary IMs (Iervolino et al., 2010). Typically, PSHA is carried out considering spectral accelerations as
43 the IM, while in the last years more efficient intensity measures have been introduced for more accurate seismic
44 structural assessment (e.g., Cordova et al., 2000; Bianchini et al., 2009; Bojorquez and Iervolino, 2011).
45 Furthermore, PSHA, as normally implemented, only refers to mainshocks (see next section) neglecting the
46 effect of foreshocks and aftershocks on seismic hazard for a site. In other words, PSHA only considers the
47 exceedance of the *im* threshold of interest due to prominent magnitude earthquakes within a cluster of events;
48 i.e., the typical way earthquakes occur (e.g., Boyd, 2012; Marzocchi and Taroni, 2014). This is to take
49 advantage of the ease of calibration and mathematical manageability of the homogeneous Poisson process
50 (HPP) (e.g., Cornell, 1968; McGuire, 2004). Nevertheless, recently, a generalized hazard integral, able to
51 account for the effect of aftershocks without losing the advantages of HPP, was developed and named
52 *sequence-based probabilistic seismic hazard analysis* or SPSHA (Iervolino et al., 2014). Finally, in some

situations, for example in the case of risk assessment of building portfolios or spatially-distributed infrastructure, in which hazard must account for exceedances at multiple sites jointly. In this case, which may be referred to as *multi-site probabilistic seismic hazard analysis* (MSPSHA), the key issue is to account for the stochastic dependence existing among the processes counting exceedances at each of the considered sites (e.g., Eguchi, 1991; Giorgio and Iervolino, 2016).

To provide an engineering-oriented tool including a number of state-of-the-art advances in probabilistic seismic hazard analysis, a stand-alone software named REgionAl, Single-SitE and Scenario-based Seismic hazard analysis (REASSESS V2.0), with a graphical user interface (GUI), has been developed and it is presented herein.¹ To this aim, the remainder of this paper is structured such that the hazard assessment methodologies considered are recalled first, along with the algorithms and numerical procedures developed for their implementation. Subsequently, REASSESS V2.0 is presented with the main input and output options. Finally, illustrative examples show the tools capabilities for earthquake engineering practice.

2. Single-site PSHA essentials

In classical PSHA, earthquakes on a seismic source are assumed to occur according to a homogeneous Poisson process (HPP) characterized by a rate, ν . In other words, the probability of observing, in the time interval ΔT , a number of earthquakes, $N(\Delta T)$, exactly equal to n is given by equation (1).

$$P[N(\Delta T) = n] = \frac{(\nu \cdot \Delta T)^n}{n!} \cdot e^{-\nu \cdot \Delta T} \quad (1)$$

The objective of PSHA is to compute the rate, λ_{im} , of seismic events exceeding the im threshold at a site of interest. Such a rate completely defines the homogeneous Poisson process (HPP) describing the occurrence of the events causing exceedance of im . In other words, the probability that, in the time interval ΔT , the number of earthquakes causing exceedance of im at the site, $N_{im}(\Delta T)$, is equal to n , is given by Equation (2).

$$P[N_{im}(\Delta T) = n] = \frac{(\lambda_{im} \cdot \Delta T)^n}{n!} \cdot e^{-\lambda_{im} \cdot \Delta T} \quad (2)$$

¹ An early release of REASSESS (V1.0) was introduced in Iervolino et al. (2016a).

75 For a site subjected to earthquakes generated at n_s seismic sources, the rate λ_{im} can be computed as illustrated
 76 in Equation (3), known as the *hazard integral*.

$$77 \quad \lambda_{im} = \sum_{i=1}^{n_s} \nu_i \cdot \int \int \int_{M \ X \ Y} P[IM > im | m, x, y]_i \cdot f_{M,XY,i}(m, x, y) \cdot dm \cdot dx \cdot dy \quad (3)$$

78 In the equation the i subscript indicates the i -th seismic source; ν_i is the rate of earthquakes above a minimum
 79 magnitude of interest ($m_{\min,i}$) and below the maximum magnitude deemed possible for the source ($m_{\max,i}$);
 80 $f_{M,XY,i}(m, x, y)$ is the joint probability density function (PDF) of earthquake magnitude M and location
 81 $\{X, Y\}$; $P[IM > im | m, x, y]_i$, typically provided by a ground motion prediction equation (GMPE), is the
 82 exceedance probability conditional on the magnitude and location (via a source-to-site distance metric).
 83 GMPEs, usually, also account for soil type, rupture mechanism and other parameters that are not explicitly
 84 considered in the notation here for the sake of simplicity (see also Section 2.1).

85 It is also only for simplicity that the location is defined in Equation (3) by means of two horizontal
 86 coordinates that can represent, for example, the epicenter. This representation is typically used in the case of
 87 areal source zones; however, it is frequent that hazard assessments have to account for three-dimensional faults
 88 (see Section 5.1). Moreover, it also happens that the distance metric of the selected GMPE is not consistent
 89 with the way location is defined. In these cases, because the relationship between location and source-to-site
 90 distance is not necessarily deterministic, the hazard integral has to account for the probabilistic distribution of
 91 the distance metric of the GMPE, conditional to the considered location parameters (e.g., Scherbaum et al.,
 92 2004).

93 Magnitude and location of the earthquake are often considered stochastically independent, that is
 94 $f_{M,XY,i}(m, x, y) = f_{M,i}(m) \cdot f_{X,Y,i}(x, y)$. The distribution $f_{M,i}(m)$ is often modeled as an exponential
 95 distribution in the $(m_{\min,i}, m_{\max,i})$ interval; i.e., of Gutenberg-Richter (G-R) type (Gutenberg and Richter, 1944);
 96 however, other choices are also considered by literature (e.g., Convertito et al., 2006). The distribution of
 97 earthquake location, $f_{X,Y,i}(x, y)$, typically reflects the hypothesis of uniformly-distributed probability on the
 98 source. For further details on classical PSHA the interested reader is referred to, for example, Reiter (1990).

Equation (3) can be numerically solved via a matrix formulation approximating the integrals with summations. To this aim, MATHWORKS-MATLAB® provides a simple computing environment that can be used to evaluate this expression. The domain of the possible realizations of the magnitude random variable (RV) is discretized via k magnitude bins represented by the values $\{m_1, m_2, \dots, m_k\}$, while the seismic source is discretized by means of s point-like seismic sources, $\{(x, y)_1, (x, y)_2, \dots, (x, y)_s\}$. Given these two vectors of size $1 \times k$ and $1 \times s$, Equation (3) can be approximated by Equation (4), where the row vector approximates $f_{x,y,i}(x, y)$ by a mass probability function (MPF) described by a vector in a way that each element is repeated k times; i.e., the first k elements are the probabilities of $(x, y)_1$, the elements from $k + 1$ until $2k$ are for $(x, y)_2$ and so on, until $(x, y)_s$. Thus, the row vector has size $1 \times (k \cdot s)$. The first column vector of Equation (4) is a $(k \cdot s) \times 1$ vector and accounts for the GMPE: each element represents the exceedance probability conditional to magnitude and location. The second column vector of the equation collects the finite k probabilities of event's magnitude, identically repeated s -times, as shown and it is, again, a $(k \cdot s) \times 1$ vector. Finally, in the equation, the pointwise multiplication between matrices of the same size (i.e., the *Hadamard product*, represented by the \otimes symbol) results in a matrix of the size of those multiplied in which each element is the product of the corresponding elements of the original matrices.

$$\lambda_{im} = \sum_{i=1}^{n_i} \nu_i \cdot \left\{ P[(x, y)_1] \quad P[(x, y)_1] \quad \cdots \quad P[(x, y)_1] \quad \cdots \quad P[(x, y)_s] \quad P[(x, y)_s] \quad \cdots \quad P[(x, y)_s] \right\}_i \cdot$$

$$\cdot \left(\begin{array}{c} P[IM > im | m_1, (x, y)_1] \\ P[IM > im | m_2, (x, y)_1] \\ \vdots \\ P[IM > im | m_k, (x, y)_1] \\ \vdots \\ P[IM > im | m_1, (x, y)_s] \\ P[IM > im | m_2, (x, y)_s] \\ \vdots \\ P[IM > im | m_k, (x, y)_s] \end{array} \right)_i \otimes \left(\begin{array}{c} P[m_1] \\ P[m_2] \\ \vdots \\ P[m_k] \\ \vdots \\ P[m_1] \\ P[m_2] \\ \vdots \\ P[m_k] \end{array} \right)_i \quad (4)$$

Equation (4), as already discussed with respect to Equation (3), is written in the case location can be defined by means of two coordinates and the distance metric of the GMPE is a deterministic function of the location.

117 Otherwise, it is necessary to account for the non-deterministic transformation of the location in source-to-site
 118 distance, which can be done in the same framework presented herein.

119 To compute the *hazard curve*, that is the function providing λ_{im} as a function of im , the hazard integral
 120 has to be computed for a number of values of im , say q in number, discretizing the domain of IM, that is
 121 $\{im_1, im_2, \dots, im_q\}$. The corresponding rates, $\{\lambda_{im_1}, \lambda_{im_2}, \dots, \lambda_{im_q}\}$, can be obtained via a single matrix operation
 122 conceptually equivalent to Equation (4); see Iervolino et al. (2016a).

123 2.1. Disaggregation

124 Disaggregation of seismic hazard (e.g., Bazzurro and Cornell, 1999) is a procedure that allows identification
 125 of the hazard contribution of one or more RVs involved in the hazard integral: e.g., magnitude and source-to-
 126 site distance, R , which, as discussed, is a function of the earthquake location. Another RV typically considered
 127 in hazard disaggregation is ε (*epsilon*). It is the number of standard deviations that $\log(im)$ is away from the
 128 median of the GMPE considered in hazard assessment. In fact, classical GMPEs are of the type in Equation
 129 (5), where $\log(im)$ is related to magnitude, distance and other parameters.

$$130 \quad \log(im) = \mu(m, r) + \theta + \sigma \cdot \varepsilon \quad (5)$$

131 In the equation, $\sigma \cdot \varepsilon$ is to a zero-mean Gaussian RV with standard deviation σ ; often it is split in *inter*- and
 132 *intra*-event components in a way that $\sigma = \sqrt{\sigma_{intra}^2 + \sigma_{inter}^2}$. The $\mu(m, r)$ term depends on magnitude and
 133 distance, θ represents one or more coefficients accounting, for example, for the soil site class. Ultimately,
 134 $\mu(m, r) + \theta$ is the mean, and the median, of the logarithms of IM given $\{m, r, \theta\}$. (Note that, although the
 135 majority of the GMPEs is of the type in Equation (5), see Stewart et al., 2015, most of the recent models have
 136 soil factors that also change with magnitude and distance. This representation is considered herein to discuss
 137 some shortcuts implemented in REASSESS and that apply only in this case; see Sections 2.3 and 4.1.)

138 The result of disaggregation is the joint PDF of $\{M, R, \varepsilon\}$ conditional to the exceedance of an IM threshold,
 139 $f_{M,R,\varepsilon|IM}$, as per Equation (6).

$$f_{M,R,\varepsilon|IM}(m,r,e) = \frac{\sum_{i=1}^{n_s} v_i \cdot I[IM > im | m,r,e] \cdot f_{M,R,\varepsilon,i}(m,r,e)}{\lambda_{im}} \quad (6)$$

In the equation, I is an indicator function that equals one if IM is larger than im for a given magnitude, distance and ε , while $f_{M,R,\varepsilon,i}(m,r,e)$ is the marginal joint PDF obtained from the product $f_{M,R,i}(m,r) \cdot f_{\varepsilon}(e)$.

From the engineering perspective, hazard disaggregation is useful to identify the characteristics of the earthquake scenarios providing the largest contribution to the hazard being disaggregated and, consequently, for hazard-consistent seismic input selection for structural assessment (e.g., Lin et al., 2013). Moreover, it is a required information to compute the conditional hazard for secondary intensity measures, which is briefly recalled in the next section. Finally, note that disaggregation can also be obtained for the occurrence of im , that is $IM = im$, and REASSESS provides also this result; i.e., McGuire (1995). For a discussion on whether exceedance or occurrence disaggregation is needed in earthquake engineering, see, for example, Fox et al. (2016).

2.2. Conditional hazard

Vector-valued probabilistic seismic hazard analysis (VPSHA), originally introduced by Bazzurro and Cornell (2002), provides the rate of earthquakes causing joint occurrence (or exceedance) of the thresholds of two IMs at the site. VPSHA could improve the accuracy in the prediction of structural damage (e.g., Baker, 2007). If one of the two intensity measures can be considered of primary importance with respect to the other, conditional hazard (Iervolino et al., 2010) can be considered an alternative to VPSHA. Conditional hazard provides the distribution of a secondary intensity measure (IM_2), conditional to the occurrence (or exceedance) of a threshold of the primary one, that is $IM_1 = im_1$ (or $IM_1 > im_1$). In the hypothesis of bivariate normality of the logarithms of the two IMs, the conditional mean ($\mu_{\log IM_2|IM_1,M,R}$) and standard deviation ($\sigma_{\log IM_2|IM_1}$) of $\log(IM_2)$, given IM_1 , magnitude and distance, are reported in Equation (7).

$$\begin{cases} \mu_{\log IM_2|IM_1,M,R} = \mu_{\log IM_2|M,R} + \rho \cdot \sigma_{\log IM_2|M,R} \cdot \frac{\log IM_1 - \mu_{\log IM_1|M,R}}{\sigma_{\log IM_1|M,R}} \\ \sigma_{\log IM_2|IM_1} = \sigma_{\log IM_2} \cdot \sqrt{1 - \rho^2} \end{cases} \quad (7)$$

162 In the equation, $\mu_{\log IM_2|M,R}$ and $\sigma_{\log IM_2|M,R}$ are the mean and standard deviation of $\log IM_2$; $\mu_{\log IM_1|M,R}$ and
 163 $\sigma_{\log IM_1|M,R}$ are the mean and standard deviation of $\log IM_1$ according to the selected GMPE (these terms are
 164 simply indicated as $\mu(m,r)$ and σ , respectively, in Section 2.1); ρ is the correlation coefficient between
 165 residuals of $\log IM_1$ and $\log IM_2$ at the site (e.g., Baker and Jayaram, 2008). Thus, the conditional distribution
 166 of the logarithm of the secondary IM is given by Equation (8) in which $f_{M,R,\varepsilon|IM_1}$ is from disaggregation and
 167 $f_{\log IM_2|IM_1,M,R,\varepsilon}$ has the parameters in Equation (7).

$$168 \quad f_{\log IM_2|IM_1}(\log im_2|im_1) = \int \int \int f_{\log IM_2|IM_1,M,R,\varepsilon}(\log im_2|im_1,m,r,e) \cdot f_{M,R,\varepsilon|IM_1}(m,r,e|im_1) \cdot dm \cdot dr \cdot d\varepsilon \quad (8)$$

169 Factually, the conditional hazard formulation of Equation (8) allows VPSHA to be an output of REASSESS.
 170 This is because, for example, $f_{\log IM_2|IM_1}$ multiplied by the absolute value of the derivative of the hazard curve
 171 from Equation (3), $|d\lambda_{im}|$, calculated in im_1 , allows to obtain the joint annual rate of $\{IM_1, IM_2\}$ for any pair
 172 of arbitrarily-selected realizations of the two IMs, $\lambda_{IM_1=im_1, IM_2=im_2}$, as per Equation (9).

$$173 \quad \lambda_{IM_1=im_1, IM_2=im_2} = |d\lambda_{im_1}| \cdot f_{\log IM_2|IM_1}(\log im_2|im_1) \quad (9)$$

174 2.3. Logic tree and shortcuts for GMPEs with additive soil factors

175 PSHA is often implemented considering a logic tree, which allows accounting for *model uncertainty* (e.g.,
 176 McGuire, 2004; Kramer, 1996); indeed, it allows the use of alternative models, each of which is assigned a
 177 weighing factor that is interpreted as the probability of that model being the *true* one. When the logic tree of
 178 n_b branches is of concern, λ_{im} is computed through Equation (10) in which p_j and $\lambda_{im,j}$ are the weight and
 179 the result of each branch of the logic tree, respectively.

$$180 \quad \lambda_{im} = \sum_{j=1}^{n_b} \lambda_{im,j} \cdot p_j \quad (10)$$

181 It should also be noted that, according to Equation (5), and only in the case of GMPEs of this type, it can be
 182 easily demonstrated that, if PSHA is performed without logic tree: (i) hazard curves for the condition
 183 represented by θ (e.g., a specific site soil class) can be obtained shifting, in the logarithmic space, those for a

reference condition when $\theta=0$; and (ii) disaggregation distribution does not depend on θ (i.e., disaggregation does not change with the soil site class). Moreover, if a logic tree featuring different GMPEs, with this same type of structures, is adopted, the discussed translation of hazard curves can be applied to the result of each branch, then re-applying Equation (10) provides the hazard in the changed conditions (Iervolino, 2016).

3. Sequence-based probabilistic seismic hazard analysis

Classical single-site PSHA discussed in the previous section neglects the effect of aftershocks on the exceedance rate. This descends from the fact that the rates ν_i , $\{1,2,\dots,n_s\}$ are obtained removing alleged foreshocks and aftershocks from earthquake catalogs; i.e., they refer to the so-called *declustered* catalogs. This is mainly because declustering is needed for the HPP to apply (Gardner and Knopoff, 1974). Recently, Boyd (2012) discussed that mainshock-aftershock sequences occur, on each seismic source, with the same rate of the mainshocks; i.e., ν_i of Equation (3). Then, Iervolino et al. (2014) demonstrated the possibility to include the effect of aftershocks in PSHA still working with HPP and declustered catalogs. On this premise, the SPSHA, was developed combining PSHA with the aftershock probabilistic seismic hazard analysis (APSHA) of Yeo and Cornell (2009). As a result, for any given *im*-value, SPSHA provides the annual rate, λ_{im} , of mainshock-aftershock sequences that cause exceedance of *im* at the site, which can be computed via Equation (11).

$$\lambda_{im} = \sum_{i=1}^{n_s} \nu_i \cdot \left\{ 1 - \int \int \int_{MXY} P[IM \leq im | m, x, y]_i \cdot e^{-E[N_{A,im|m}(0, \Delta T_A)]} \cdot f_{M,X,Y,i}(m, x, y) \cdot dm \cdot dx \cdot dy \right\} \quad (11)$$

In the equation, the terms: ν_i , $P[IM \leq im | m, x, y]_i = 1 - P[IM > im | m, x, y]_i$, and $f_{M,X,Y,i}(m, x, y)$ are the same defined in Equation (3). The exponent $E[N_{A,im|m}(0, \Delta T_A)]$ refers to aftershocks, as indicated by the *A* subscript: it represents the average number of aftershocks that cause exceedance of *im* in a sequence triggered by the mainshock of magnitude and location $\{m, x, y\}$, Equation (12).

$$\begin{aligned} E[N_{A,im|m}(0, \Delta T_A)] &= \\ &= E[N_{A|im}(0, \Delta T_A)] \cdot \int \int \int_{M_A X_A Y_A} P[IM > im | m_A, x_A, y_A]_i \cdot f_{M_A, X_A, Y_A, i | M, X, Y}(m_A, x_A, y_A | m, x, y) \cdot dm_A \cdot dx_A \cdot dy_A \end{aligned} \quad (12)$$

207 The probability represented by the exponential term depends on $P[IM > im | m_A, x_A, y_A]_i$, that is the probability
 208 that im is exceeded given an aftershock of magnitude and location identified by the vector $\{m_A, x_A, y_A\}$; i.e.,
 209 a GMPE for aftershocks (although in several applications those for mainshock are considered applicable). The
 210 term $f_{M_A, X_A, Y_A, i | M, X, Y}$ is the distribution of magnitude and location of aftershocks, which is conditional on the
 211 features, $\{m, x, y\}$, of the mainshock. This distribution can be written as $f_{M_A, X_A, Y_A, i | M, X, Y} = f_{M_A, i | M} \cdot f_{X_A, Y_A, i | M, X, Y}$,
 212 where $f_{M_A, i | M}$ is the PDF of aftershock magnitude of G-R type, and $f_{X_A, Y_A, i | M, X, Y}$ is the distribution of the
 213 location of the aftershocks and depends on the magnitude and location of the mainshock (e.g., Utsu, 1970).
 214 $E[N_{A|m}(0, \Delta T_A)]$ is the expected number of aftershocks to the mainshock of magnitude m in the ΔT_A and,
 215 according to Yeo and Cornell (2009), can be computed via Equation (13) in which $m_{A,min}$ is the minimum
 216 magnitude considered for aftershocks (often taken equal to the minimum magnitude considered for
 217 mainshocks) and $\{a, b, c, p\}$ are parameters of the *modified Omori Law*.

$$218 \quad E[N_{A|m}(0, \Delta T_A)] = \frac{10^{a+b(m-m_{A,min})} - 10^a}{p-1} \cdot \left[c^{1-p} - (\Delta T_A + c)^{1-p} \right] \quad (13)$$

219 Finally, note that λ_{im} is still the rate of the HPP of the kind in Equation (2), which now regulates the occurrence
 220 of sequences causing exceedance of im .

221 The matrix formulation presented in Equation (4) for the numerical computation of PSHA, can be extended
 222 to the SPSHA case as reported in Equation (14). In the latter, vectors are arranged as discussed referring to
 223 Equation (4), but a new column vector is introduced: it has the same $(k \cdot s) \times 1$ size and each element of it
 224 accounts for the probability that none of the aftershocks, to the mainshock of given magnitude and location,
 225 cause the exceedance of im .

$$\begin{aligned}
\lambda_{im} = & \sum_{i=1}^{n_s} \nu_i \cdot \left[1 - \left\{ P[(x, y)_1] \quad P[(x, y)_1] \quad \cdots \quad P[(x, y)_1] \quad \cdots \quad P[(x, y)_s] \quad P[(x, y)_s] \quad \cdots \quad P[(x, y)_s] \right\}_i \cdot \right. \\
& \left. \left\{ \begin{array}{c} P[IM \leq im \mid m_1, (x, y)_1] \\ P[IM \leq im \mid m_2, (x, y)_1] \\ \vdots \\ P[IM \leq im \mid m_k, (x, y)_1] \\ \vdots \\ P[IM \leq im \mid m_1, (x, y)_s] \\ P[IM \leq im \mid m_2, (x, y)_s] \\ \vdots \\ P[IM \leq im \mid m_k, (x, y)_s] \end{array} \right\}_i \otimes \left\{ \begin{array}{c} e^{-E[N_{A|m_1}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_1, (x, y)_1]} \\ e^{-E[N_{A|m_2}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_2, (x, y)_1]} \\ \vdots \\ e^{-E[N_{A|m_k}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_k, (x, y)_1]} \\ \vdots \\ e^{-E[N_{A|m_1}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_1, (x, y)_s]} \\ e^{-E[N_{A|m_2}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_2, (x, y)_s]} \\ \vdots \\ e^{-E[N_{A|m_k}(0, \Delta T_A)] \cdot P[IM_A > im \mid m_k, (x, y)_s]} \end{array} \right\}_i \otimes \left\{ \begin{array}{c} P[m_1] \\ P[m_2] \\ \vdots \\ P[m_k] \\ \vdots \\ P[m_1] \\ P[m_2] \\ \vdots \\ P[m_k] \end{array} \right\}_i \right] \quad (14)
\end{aligned}$$

3.1. SPSHA disaggregation

Disaggregation of seismic hazard can be performed also in the case of SPSHA. Equation (15) provides the PDF of mainshock magnitude and distance (R), given that the ground motion intensity of the mainshock, IM , or the maximum ground motion intensity of the following aftershock sequence (IM_{UA}) is larger than the im threshold. In the equation, similarly to what discussed in Section 2.1, $\{X, Y\}$ and $\{X_A, Y_A\}$ vector RVs, are substituted by R and R_A , respectively.

$$\begin{aligned}
f_{M,R|IM > im \cup IM_{UA} > im}(m, r) = & \frac{\sum_{i=1}^{n_s} \nu_i}{\lambda_{im}} \times \\
& \times \left\{ 1 - P[IM \leq im \mid m, r]_i \cdot e^{-E[N_{A|im}(0, \Delta T_A)] \cdot \int_{M_A, R_A} P[IM_A > im \mid m_A, r_A]_i \cdot f_{M_A, R_A|M, R, i}(m_A, r_A \mid m, r) \cdot dm_A \cdot dr_A} \right\} \cdot f_{M, R, i}(m, r) \quad (15)
\end{aligned}$$

Moreover, it can be useful to quantify the probability that, given the im exceedance, such exceedance is caused by an aftershock rather than by a mainshock. This probability, which quantifies the contribution of aftershocks to hazard, is recalled in Equation (16).

$$\begin{aligned}
P[IM_{UA} > im \cap IM \leq im \mid IM > im \cup IM_{UA} > im] = & \sum_{i=1}^{n_s} \frac{\nu_i}{\lambda_{im}} \cdot \iint_{M, R} P[IM \leq im \mid m, r]_i \times \\
& \times \left(1 - e^{-E[N_{A|im}(0, \Delta T_A)] \cdot \int_{M_A, R_A} P[IM_A > im \mid m_A, r_A]_i \cdot f_{M_A, R_A|M, R, i}(m_A, r_A \mid m, r) \cdot dm_A \cdot dr_A} \right) \cdot f_{M, R, i}(m, r) \cdot dm \cdot dr \quad (16)
\end{aligned}$$

238 In the equation, $P[IM_{\cup A} > im \cap IM \leq im | IM > im \cup IM_{\cup A} > im]$ it is the probability that, given that
 239 exceedance of im has been observed during the mainshock-aftershock sequence, $(IM > im \cup IM_{\cup A} > im)$, it
 240 was in fact an aftershock to cause it, while the mainshock was below the threshold; i.e., $(IM_{\cup A} > im \cap IM \leq im)$
 241 . All the terms of the equation have been already defined discussing Equation (11); see Iervolino et al. (2018)
 242 for derivation of the equation.

243 4. Multi-site hazard

244 In the case of MSPSHA, for a set of spatially-distributed sites, say n_{sts} in number, one can define a vector of
 245 thresholds, one for each site $\{im_1, im_2, \dots, im_{n_{sts}}\}$, of the IM of interest. Given a vector of thresholds, the sought
 246 outcomes of MSPSHA can be various, for example, probabilistic distribution of the total number of
 247 exceedances collectively observed at the sites in the ΔT time interval. The main issue with MSPSHA is that,
 248 even if the process counting exceedance at each of the sites is an HPP, that is Equation (2), these HPPs are (in
 249 general) not independent. Then, the process that counts the total number of exceedances observed at the
 250 ensemble of the sites over time is not a HPP. The nature and form of stochastic dependence, existing among
 251 the processes counting in time exceedances of ground motion thresholds at multiple sites, is related to the
 252 probabilistic characterization of the effects of a common earthquake at the different sites (e.g., Giorgio and
 253 Iervolino, 2016).

254 The same reasoning discussed for one IM at multiple sites, can be applied when MPSHA involves multiple
 255 IMs. For example, if one considers as IMs two pseudo accelerations at two spectral periods, $IM_1 = Sa(T_1)$ and
 256 $IM_2 = Sa(T_2)$, it is generally assumed that, given an earthquake of m and $\{x, y\}$ characteristics, the logarithms
 257 of IMs at the sites form a Gaussian random field (GRF), a realization of which is a $1 \times (n_{sts} \cdot 2)$ vector of the
 258 type $\{im_{1,1}, im_{1,2}, \dots, im_{1,n_{sts}}, im_{2,1}, im_{2,2}, \dots, im_{2,n_{sts}}\}$. This means that the logarithms of IMs have a multivariate
 259 normal distribution, where the components of the mean vector are given by the $E[\log IM_1 | m, r_j, \theta]$ and
 260 $E[\log IM_2 | m, r_j, \theta]$ terms; two for each j , being r_j the distance between the site j and the location of the

261 seismic event, and the covariance matrix, Σ , is given in Equation (17). In the equation, $\sigma_{\text{inter},1}$ and $\sigma_{\text{inter},2}$ are
 262 the standard deviations of the inter-event residuals of the GMPEs of the two IMs, while $\sigma_{\text{intra},1}$ and $\sigma_{\text{intra},2}$ are
 263 the standard deviation of intra-event residuals of $Sa(T_1)$ and $Sa(T_2)$, respectively; $\rho_{\text{inter}}(T_1, T_2)$ is the
 264 correlation coefficient between inter-event residuals at the two spectral periods in the same earthquake, while
 265 $\rho_{\text{intra}}(T_1, T_2, h_{i,j})$ is the correlation coefficient between intra-event residuals of the GMPEs of $Sa(T_1)$ and
 266 $Sa(T_2)$ for sites i and j ; and $h_{i,j}$ is the inter-site distance. In this case, Σ is the sum of two square matrices,
 267 each of $(n_{\text{sts}} \cdot 2) \times (n_{\text{sts}} \cdot 2)$ size. The first matrix accounts for the correlation of inter-event residuals which
 268 is, by definition, independent on the inter-site distance; the second matrix accounts for the intra-event residuals
 269 correlation and is dependent on inter-site distance as well as the selected spectral periods. Assigning the mean
 270 vector and the covariance matrix completely defines the GRF in one earthquake (e.g., Baker and Jayaram,
 271 2008; Esposito and Iervolino, 2012; Loth and Baker, 2013; Markhvida et al., 2017).

$$\begin{aligned}
 \Sigma = & \begin{bmatrix} \sigma_{\text{inter},1}^2 & \cdots & \sigma_{\text{inter},1}^2 & \rho_{\text{inter}}(T_1, T_2) \cdot \sigma_{\text{inter},1} \cdot \sigma_{\text{inter},2} & \cdots & \rho_{\text{inter}}(T_1, T_2) \cdot \sigma_{\text{inter},1} \cdot \sigma_{\text{inter},2} \\ & \ddots & \vdots & \vdots & \ddots & \vdots \\ & & \sigma_{\text{inter},1}^2 & \rho_{\text{inter}}(T_1, T_2) \cdot \sigma_{\text{inter},1} \cdot \sigma_{\text{inter},2} & \cdots & \rho_{\text{inter}}(T_1, T_2) \cdot \sigma_{\text{inter},1} \cdot \sigma_{\text{inter},2} \\ & & & \sigma_{\text{inter},2}^2 & \cdots & \sigma_{\text{inter},2}^2 \\ & & \text{sym} & & \ddots & \vdots \\ & & & & & \sigma_{\text{inter},2}^2 \end{bmatrix} + \\
 & + \begin{bmatrix} \sigma_{\text{intra},1}^2 & \cdots & \rho_{\text{intra}}(T_1, T_1, h_{1,n_{\text{sts}}}) \cdot \sigma_{\text{intra},1}^2 & \rho_{\text{intra}}(T_1, T_2, h_{1,1}) \cdot \sigma_{\text{intra},1} \cdot \sigma_{\text{intra},2} & \cdots & \rho_{\text{intra}}(T_1, T_2, h_{1,n_{\text{sts}}}) \cdot \sigma_{\text{intra},1} \cdot \sigma_{\text{intra},2} \\ & \ddots & \vdots & \vdots & \ddots & \vdots \\ & & \sigma_{\text{intra},1}^2 & \rho_{\text{intra}}(T_1, T_2, h_{n_{\text{sts}},1}) \cdot \sigma_{\text{intra},1} \cdot \sigma_{\text{intra},2} & \cdots & \rho_{\text{intra}}(T_1, T_2, h_{n_{\text{sts}},n_{\text{sts}}}) \cdot \sigma_{\text{intra},1} \cdot \sigma_{\text{intra},2} \\ & & & \sigma_{\text{intra},2}^2 & \cdots & \rho_{\text{intra}}(T_2, T_2, h_{1,n_{\text{sts}}}) \cdot \sigma_{\text{intra},2}^2 \\ & & \text{sym} & & \ddots & \vdots \\ & & & & & \sigma_{\text{intra},2}^2 \end{bmatrix}
 \end{aligned}
 \tag{17}$$

274 To compute MPSHA representing the GRF with the discussed covariance structure, in REASSESS the Monte
 275 Carlo simulation approach has been chosen. In this framework, one possible algorithm is the two-step
 276 procedure of Figure 1, which is described, for simplicity, with reference to a single seismic source where
 277 earthquakes occur as per Equation (1) with assigned magnitude and location distributions.

(a) The first step is addressed to simulate and collect realizations of the GRF conditional to the occurrence of an earthquake of generic magnitude and location. In other words, magnitudes and locations of the seismic events on the source are sampled according to their distributions and, then, the realizations of the IMs at the considered sites are simulated in accordance with the considered GMPEs and Σ . This step is described in Figure 1, where n_m , n_{xy} and n_ε are the indices counting the number of simulations for magnitude, event location and GRF of residuals at the sites, respectively; capital letters of the indices, N_m , N_{xy} and N_ε are the total number of simulations for each of the three variables. Thus, the results of the first step are $N_m \cdot N_{xy} \cdot N_\varepsilon$ vectors, one for each simulation, collecting the IM-values simulated at the sites in each event. Each vector $\{\underline{im}\} = \{im_1, im_2, \dots, im_{n_{sts}}\}$ represents realizations of the random field of IMs at the sites in one generic (i.e., considering all possible magnitudes and locations) earthquake and, therefore, it is time-invariant.

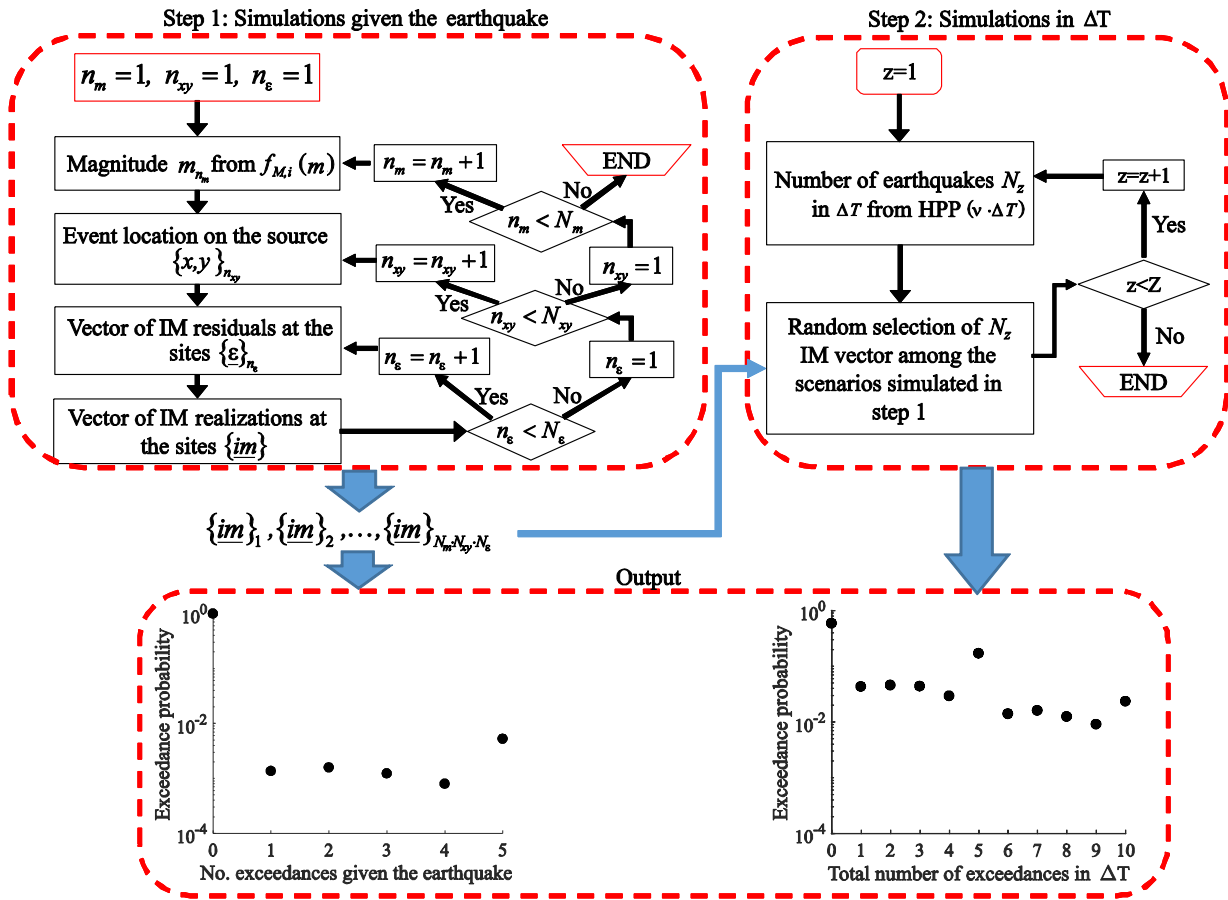


Figure 1. Flowchart of the simulation procedure for MSPSHA in the case of single seismic source.

(b) The realizations from step (a) are the input for this step that consists of simulating the process of earthquakes affecting sites, in any time interval ΔT of interest; i.e., the *seismic history for the sites* in ΔT . In each run of the simulation of this step, indicated by the index z which varies from 1 to Z , the number of earthquake events on the source is sampled from a HPP with mean equal to $\nu \cdot \Delta T$. Then, a number of IM random fields, equal to the sampled number of events, is randomly selected among those generated in the first step of the procedure. These random fields collectively represent one realization of the seismic history at the sites in ΔT . Therefore, repeating Z times this step, can provide a sample of histories of what could occur in ΔT at the sites.

The simulated seismic histories can be used to compute any MSPSHA result. For example, if one is interested in the distribution (i.e., the MPF) of the total number of exceedances collectively observed at the sites in ΔT , it is sufficient to count in how many histories a specific number of total exceedances of the $\{im_1, im_2, \dots, im_{n_{st}}\}$ vector has been observed and divide by the total number of simulated histories. For example, the probability that zero exceedances are observed collectively at the site, in ΔT years, is equal to the number of histories in which none of the IM thresholds set for each of the sites is exceeded, divided by the number of simulated histories (i.e., Z).

In the case of more than one seismic source, the first step is repeated for each of them to simulate the random field they individually produce. In the second step, the HPP describing the event occurrence on all the sources has mean equal to $\Delta T \cdot \sum_i \nu_i$. This, similarly to the case of a single source, is used to sample the number of earthquakes in ΔT and to randomly select the random field realizations from those of each source; the number of realizations to be selected for each source is proportional to the probability that given that an earthquake occurs it is from source i , that is $\nu_i / \sum_i \nu_i$. At this point the seismic history in ΔT for the sites is obtained in analogy to the case of a single source.

4.1. MSPSHA shortcuts for GMPEs with additive factors

In this section some helpful shortcuts for MSPSHA calculations that are implemented in REASSESS and that apply (only) in the case of GMPEs of the type in Equation (5) are discussed. It should be noted that the covariance of two or more RVs does not change adding constant terms. Thus, to the aim of this section, it is

required to recognize that Equation (5) implies that the RV representing the logarithms of IM for a site with conditions represented by θ , is obtained adding such a coefficient to the RV representing the logarithms of IM for a reference condition for which $\theta=0$; this means that the covariance matrix of the GRF is also independent of θ (e.g., the soil class of each of the sites). As a consequence, the simulations described in Section 4 can be carried out considering a common site condition for all sites (e.g., rock). To obtain GRF realizations reflecting the different site conditions at the sites from those for the reference case, it is sufficient to add to the logarithms of the simulated IMs the site-specific coefficient, that is $\{\theta_1, \theta_2, \dots, \theta_{n_{sts}}\}$, from the GMPE. Equivalent, but even simpler, is to subtract the $\{\theta_1, \theta_2, \dots, \theta_{n_{sts}}\}$ vector from the vectors of logarithms of the IM thresholds for the sites. However, in closing this section, it has to be emphasized that, as mentioned, several recent GMPEs are not of the type in Equation (5) for what concerns the soil term, and these shortcuts do not apply (see also Stafford et al., 2017). Nevertheless, this same reasoning holds in the case θ of Equation (5) represents any other factor affecting the IMs, not only soil site class.

5. REASSESS V2.0

To implement the types of hazard assessment discussed above, REASSESS V2.0 is coded in MATLAB and profits of a graphical user interface (GUI). The GUI features one input panel and two output panels, one for PSHA/SPSHA and one for MSPSHA. In fact, the main GUI is complemented by secondary interfaces that pop up when needed (see Figure 2). Note that, in the case of extended analyses (e.g., several seismic sources or sites), input can also be defined via dedicated MICROSOFT®-EXCEL spreadsheets, as a shortcut.

A schematic flowchart of the way REASSESS V2.0 operates is given in Figure 3. First, the user is required to define the type of analysis to be performed; i.e., PSHA, SPSHA, or MSPSHA. Even in the case of single-site analysis (PSHA and SPSHA) the user is allowed to define more than one site of interest; in this case, REASSESS will run single-site PSHA or SPSHA separately for each of them according to Section 2 or Section 3. If MSPSHA is selected, more than one site must be defined, and the analyses are performed according to what discussed in Section 3.1. (When SPSHA or MSPSHA is selected, the corresponding PSHA is also performed for the considered sites, as it is considered a reference case.)

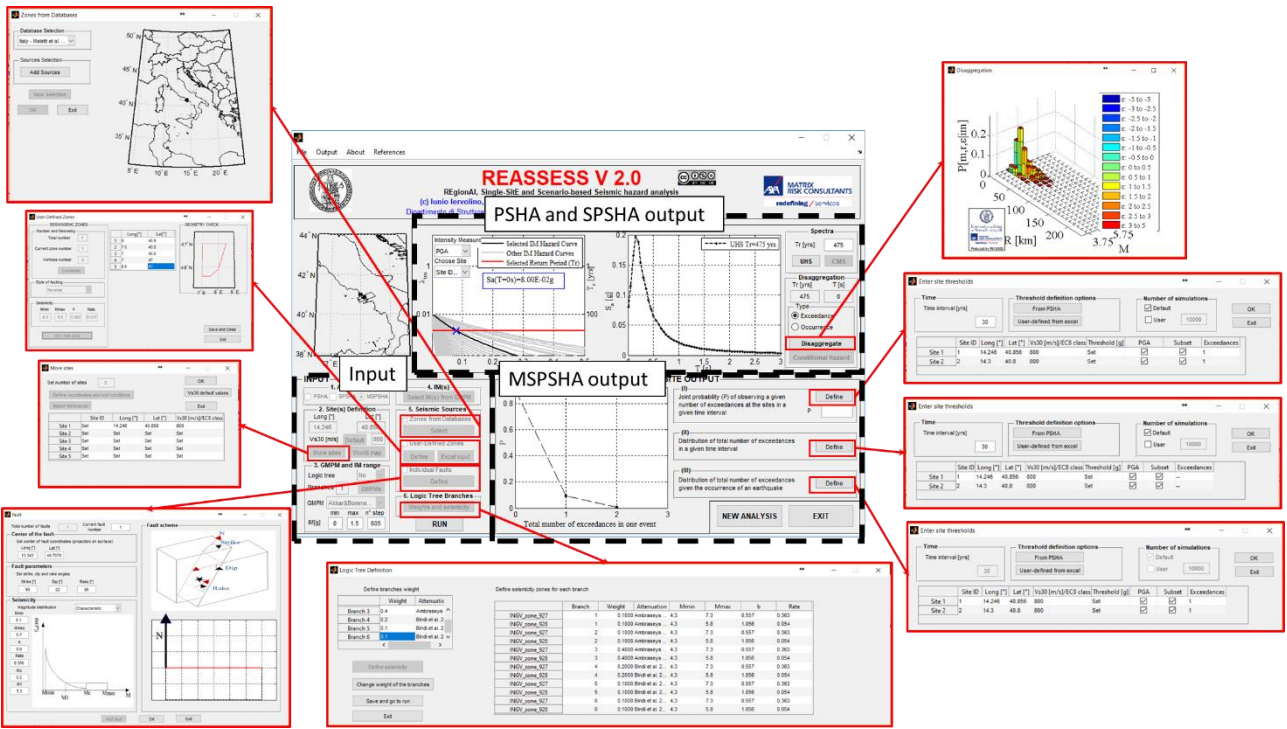


Figure 2. Principal and auxiliary GUIs of REASSESS V2.0.

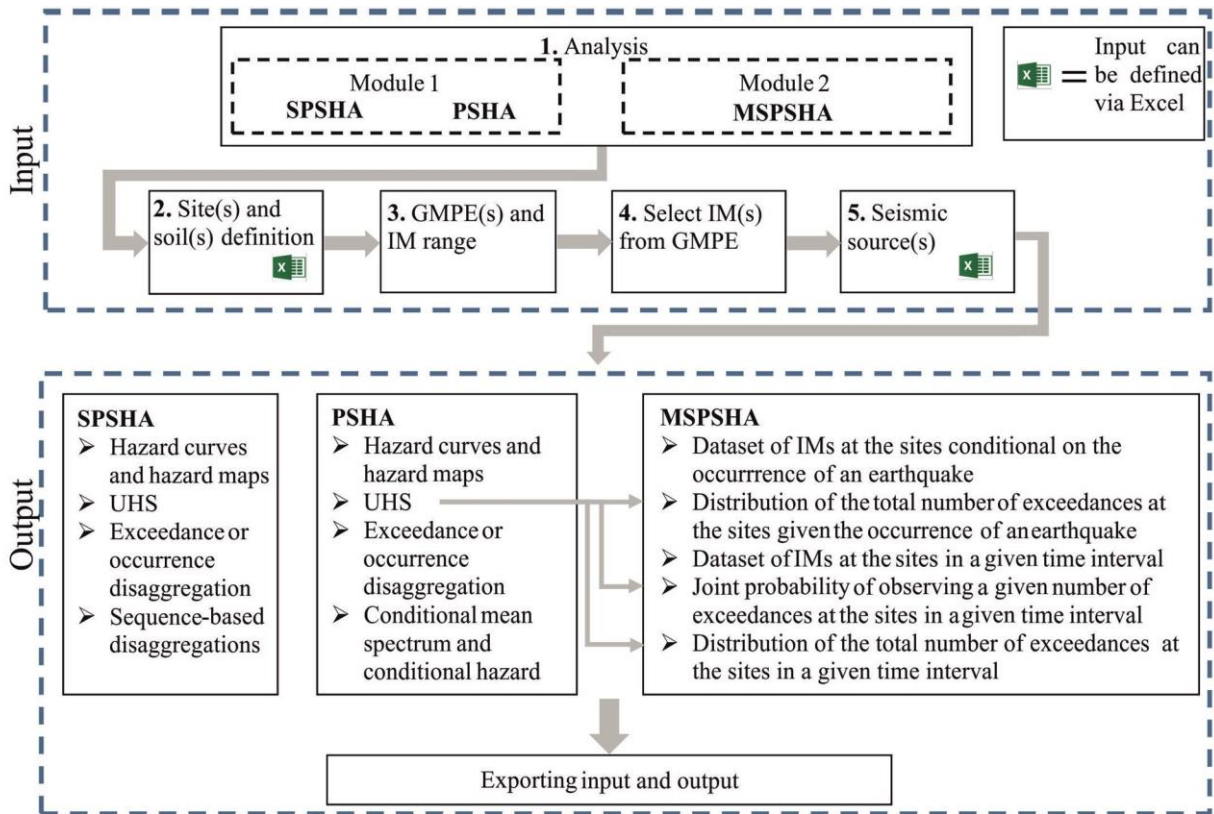


Figure 3. REASSESS V2.0 flowchart showing single-site and multisite modules functionalities.

The second step refers to definition of the coordinates and soil condition of the sites. It can be carried out via the GUI or via an EXCEL spreadsheet, for which a template is given. The soil conditions can be defined in

348 terms of shear wave velocity of the top 30 meters of subsurface profile (Vs30) expressed in meter/second or
349 in terms of the soil classes (A, B, C, D and E) according to the Eurocode 8 classification of sites (CEN, 2004).

350 The third step is dedicated to the selection of the GMPE(s). A database of alternative GMPEs is included
351 in the current release of REASSESS: Ambraseys et al. (1996), fitted on a European dataset, Akkar and Bommer
352 (2010), which refers to data from southern Europe, North Africa, and active areas of the Middle East, Bindi et
353 al. (2011), fitted on Italian dataset and Cauzzi et al. (2015), based on a worldwide dataset.² At this step, also
354 the discretization of the domain of the intensity measure for single-site PSHA, which serves to lump the hazard
355 curves, has to be defined in terms of minimum, maximum values and number of intermediate steps (constant
356 in logarithmic scale). In the case of PSHA, the third step also allows the definition of a logic tree (section 2.3)
357 in terms of: (i) parameters of the magnitude distributions, (ii) mean annual frequency of earthquake occurrence
358 on the sources and (iii) GMPEs (among those available).

359 The choice of the IMs to be considered (e.g., spectral pseudo-acceleration for different natural vibration
360 periods) for all the types of analysis (PSHA, SPSHA or MSPSHA) is dependent on the IMs available per the
361 selected GMPE (step 4). If a logic tree with different GMPE for each branch has been defined, the selection is
362 among the IMs of the GMPE belonging to the branch with the highest weight. If different branches have the
363 same weight, the selection is among the IMs of the GMPE selected for the first branch.

364 When PSHA is of concern, REASSESS also allows to perform analysis for advanced spectral-shape-based
365 intensity measures such as I_{Np} proposed by Bojórquez and Iervolino (2011) and reported in Equation (18) in
366 logarithmic. The I_{Np} is a proxy of the pseudo-acceleration response (Sa) spectral shape in a range of periods
367 $(T_1...T_N)$ and is dependent on a reference period (\bar{T}) belonging to the $(T_1...T_N)$ interval and an α parameter.
368 In its analytical expression $Sa_{avg}(T_1...T_n)$ appears; it is the geometric mean of the spectral acceleration in the
369 $(T_1...T_N)$ range of periods (Baker and Cornell, 2006a). In the software, $(T_1...T_N)$, \bar{T} and α can be selected by
370 the user (the periods can be chosen among those of the selected GMPE). It is easy to see that when the α
371 parameter equals one, I_{Np} corresponds to $Sa_{avg}(T_1...T_n)$.

² These GMPEs are of the type in Equation (5), then the shortcuts discussed in Section 2.3 and Section 4.1 apply. Also note that although the Ambraseys et al. (1996) GMPEs dates more than twenty years ago, it has been considered because it is the one the current official Italian hazard model is based on (Stucchi et al., 2011).

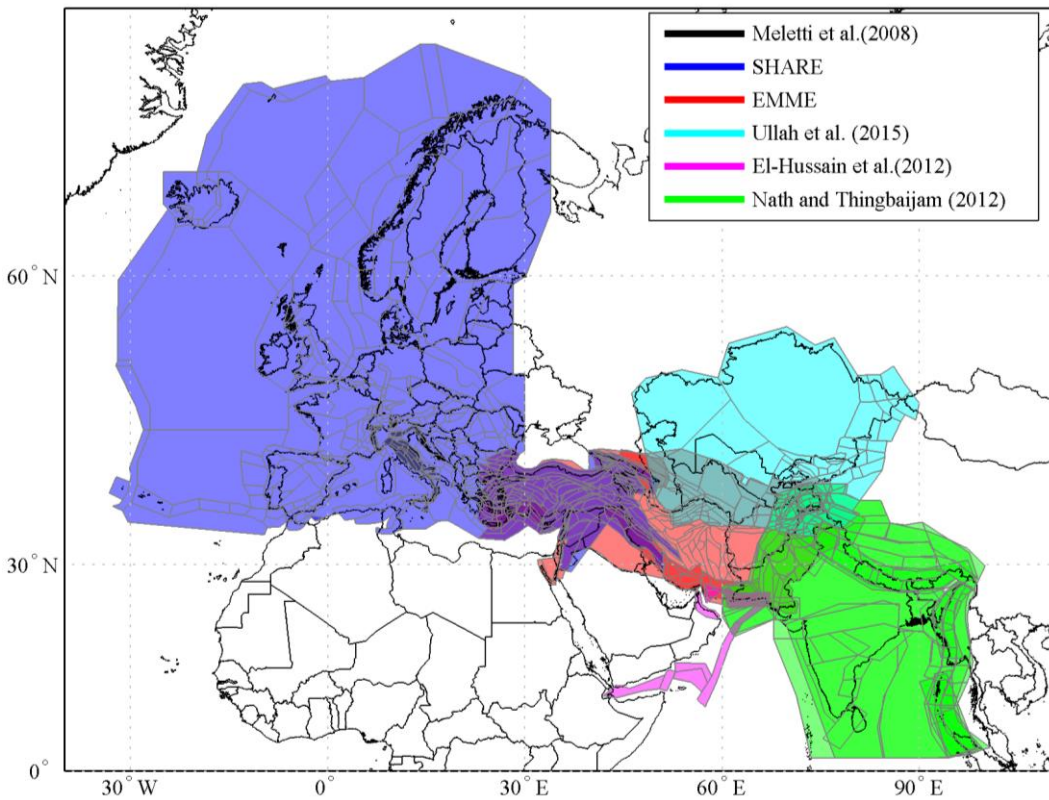
$$\log(I_{Np}) = \log[Sa(\bar{T})] + \alpha \log \left[\frac{Sa_{avg}(T_1 \dots T_N)}{Sa(\bar{T})} \right] \quad (18)$$

In the case of MSPSHA, when a single spectral ordinate is selected as IM, the user is allowed to choose the model of spatial correlation of intra-event residuals of Esposito and Iervolino (2012) or Loth and Baker (2013). On the other hand, when the IMs at the sites are spectral ordinates for several natural vibration periods, simulated cross-correlated scenarios are computed adopting the models of (i) Loth and Baker (2013) for the spatial correlation of intra-event residuals and (ii) Baker and Jayaram (2008) for the spectral correlation of inter-event residuals.

Step 5 is dedicated to the seismic source definition. In REASSESS V2.0, seismic source zones and/or finite three-dimensional faults can both be input of analysis. Faults are discussed in section 5.1, for what concerns source zones, these are defined by the coordinates of the vertices of the zone, the annual of rate of occurrence of earthquakes of Equation (1) and the event's magnitude distribution, which is assumed to be a truncated exponential distribution as discussed in Section 2; hence, the slope of the G-R relationship, together with minimum and maximum values of magnitude, is required. If known, a rupture faulting style can be associated to the seismic zone. As mentioned, all the required parameters can be alternatively given via GUI or EXCEL spreadsheet.

A number of literature databases of seismic zones are already embedded in the current version of the software. Referring to Italy, it is known that the seismic hazard study of Stucchi et al. (2011) lies at the basis of the hazard assessment for the Italian current building code and features a logic tree made of several branches; the branch named 921 is the one producing the results claimed to be the closest to those provided by the full logic tree. This branch considers the seismic source model of thirty-six areal zones of Meletti et al. (2008) and the GMPE by Ambraseys et al. (1996). It is implemented in REASSESS V2.0 and is named *Meletti et al. (2008) – Magnitude rates from DPC-INGV-S1 - Branch 921*. It is the sole database selection which implies a specific GMPE (automatically selected). An alternative source model for Italy is named *Meletti et al. (2008) – Magnitude rates from Barani et al. (2009)* in which the same source model of Meletti et al. (2008) is considered, but the associated seismic characterization is from Barani et al. (2009). Other databases in REASSES are the one from the SHARE project, which covers the Euro-Mediterranean region, the one from the EMME project, which covers middle-east; i.e., Afghanistan, Armenia, Azerbaijan, Cyprus, Georgia, Iran,

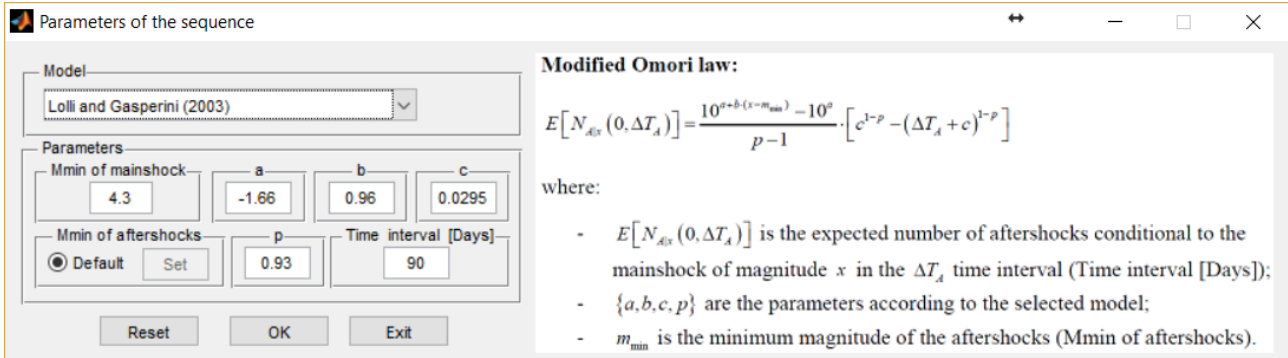
399 Jordan, Lebanon, Pakistan, Syria and Turkey. Moreover, included databases are: El-Hussain et al. (2012),
 400 Ullah et al. (2015) and Nath and Thingbaijam (2012), referring to the Sultanate of Oman, Kazakhstan,
 401 Kyrgyzstan, Tajikistan, Uzbekistan and Turkmenistan, and India, respectively. The area covered by the
 402 embedded databases is given in Figure 4. For each of the cited databases, assuming a uniform earthquake
 403 location distribution in each seismic source, epicentral distance is converted into the metric required by the
 404 GMPE according to Montaldo et al. (2005). The style-of-faulting correction factors proposed by Bommer at
 405 al. (2003) are also applied to the GMPE in accordance with the rupture mechanism associated to each seismic
 406 zone (if any).



407
 408 Figure 4. Embedded databases of seismogenic sources.

409 When SPSHA is performed, an additional step is required in the input definition. In particular, the model
 410 describing the aftershock occurrence has to be specified, that is the parameters of Equation (13), providing the
 411 expected number of aftershock in any time interval given the magnitude of the mainshock. The available
 412 models are those of Reasenber and Jones (1989 and 1994), Lolli and Gasperini (2003) and Eberhart-Phillips
 413 (1998) which refer to generic California, Italian and New Zealand aftershock sequence, respectively. Such
 414 models, can be selected through a dedicated window (Figure 5), automatically opened by REASSESS before

415 running the SPSHA. In the current version of the software, the GMPE selected for PSHA is also applied to
 416 account for the evaluation of aftershock's IM.



417
 418 Figure 5. Graphical interface window for calibration of the aftershock occurrence models.

419 5.1. Finite faults

420 REASSESS also allows to compute hazard analysis (both PSHA and MSPSHA) in the case the seismic sources
 421 are represented by means of one or more finite faults. There are many alternative ways to define the
 422 characteristics of a fault for hazard assessment purposes (Scherbaum et al., 2004). In the current version of
 423 REASSESS a fault is defined by means of a point representing its center and the dip, rake, and strikes angles
 424 (Aki and Richard, 1980). In the case of a finite fault in REASSESS, PSHA is carried out according to Equation
 425 (19), which is an adaptation of Equation (3).

$$426 \quad \lambda_{im} = \nu \cdot \int \int \int \int P[IM > im|m, x, y] \cdot f_{S|A}(s|a) \cdot f_{A|M}(a|m) \cdot f_M(m) \cdot f_{X,Y}(x, y) \cdot ds \cdot da \cdot dm \cdot dx \cdot dy \quad (19)$$

427 In the equation: ν is the rate; $\{x, y\}$ is the position of the center of the rupture with respect to the center of the
 428 fault and its distribution $f_{X,Y}(x, y)$ is taken according to Mai et al. (2005); $f_M(m)$ is the magnitude
 429 distribution that can be defined as G-R or characteristic (e.g., Convertito et al., 2006); $f_{A|M}(a|m)$ is the
 430 distribution of the rupture size, conditional to the magnitude that is modelled according to Wells and
 431 Coppersmith (1994); finally, $f_{S|A}(s|a)$ is the aspect ratio (length-to-width ratio) of the rupture and is
 432 probabilistically modeled lognormally according to Iervolino et al. (2016b).³

³ The depth of the top of the rupture is assumed to be equal to five kilometers for all events of magnitude less than 6.5 and one kilometer for events of larger magnitude, following the practice of the U.S. Geological Survey; however, this constraint is not strictly needed and could be relaxed in updated versions of REASSESS.

433 6. Output of the analyses

434 At the end of the analysis, the outputs provided by the software can be consulted via the GUI in the format of
435 figure or text files. Moreover, a compressed folder with all the input and output (figures and files) of the
436 analyses can be saved by the user. In the following sub-sections, the available results are described.

437 6.1. PSHA and SPSHA results

438 When the analysis is finished, the hazard curves are plotted in the single-site output panel (see Figure 2). If the
439 analysis is performed for more than one site, the curves for each site of interest can be selected (via a dropdown
440 menu). The uniform hazard spectrum (UHS) can be computed, and plotted in a dedicated panel, selecting any
441 return period available (that depends on the range of IMs defined at the beginning).

442 In addition, REASSESS is able to provide disaggregation of PSHA, conditional mean spectrum (CMS; Lin
443 et al., 2013) and conditional hazard (see Section 2.1 and 2.2). The conditional hazard can be computed by
444 REASSESS V2.0 profiting of the model of Bradley (2012), which provides correlation between peak ground
445 velocity (PGV) and spectral accelerations and the model of Baker and Jayaram (2008), which provides the
446 correlation among spectral acceleration values at different spectral periods. Therefore, the distribution of PGV
447 or pseudo-acceleration response spectra at any vibration period conditional to the occurrence of any spectral
448 ordinate can be computed.

449 Results of SPSHA are similar to those for PSHA; however, disaggregation is of two kinds (see Section 3.1).
450 The first is the joint probability density function of magnitude and distance of the mainshock conditional to
451 the exceedance, or the occurrence, of a chosen hazard threshold during the corresponding cluster (mainshock
452 and subsequent aftershocks). This is equivalent to the classical hazard disaggregation, in terms of magnitude
453 and distance, but computed in accordance with the approach of the SPSHA, Equation (15). The second
454 disaggregation provided represents the probability that, given that exceedance of im has been observed during
455 the mainshock-aftershock sequence, it was in fact an aftershock to cause it, Equation (16).

456 6.2. MSPSHA results

457 MSPSHA can be performed on all or on a subset of the sites defined at the beginning of the analysis. It is
458 performed through the two-steps procedure described in Section 4. At the end of the first step, the simulated
459 scenarios of IM realizations at the sites, given the occurrence of an earthquake on the sources are available. As

460 a reference, these results are also used to first provide single-site PSHA as per Equation (2) (in fact, single-site
461 PSHA can be viewed as a special case of MSPSHA; Giorgio and Iervolino, 2016) and the results are reported
462 in the single-site panel. Specifically referring to MSPSHA, REASSESS V2.0 provides three kinds of results:

- 463 (i) the probability of observing an arbitrarily chosen number of exceedances at the sites in a given
464 time interval;
- 465 (ii) the distribution of the total number of exceedances at the sites in a given time interval;
- 466 (iii) the distribution of the number of exceedances at the sites given the occurrence of an earthquake
467 (a time-invariant results).

468 Results (i) and (ii) are computed by REASSESS for any time interval without repeating the simulations of the
469 first step of analysis thus reducing the required time of computation. Text files with the GRFs simulated (i.e.,
470 realizations) conditional to a generic event and in the selected time interval are also available at the end of the
471 analyses.

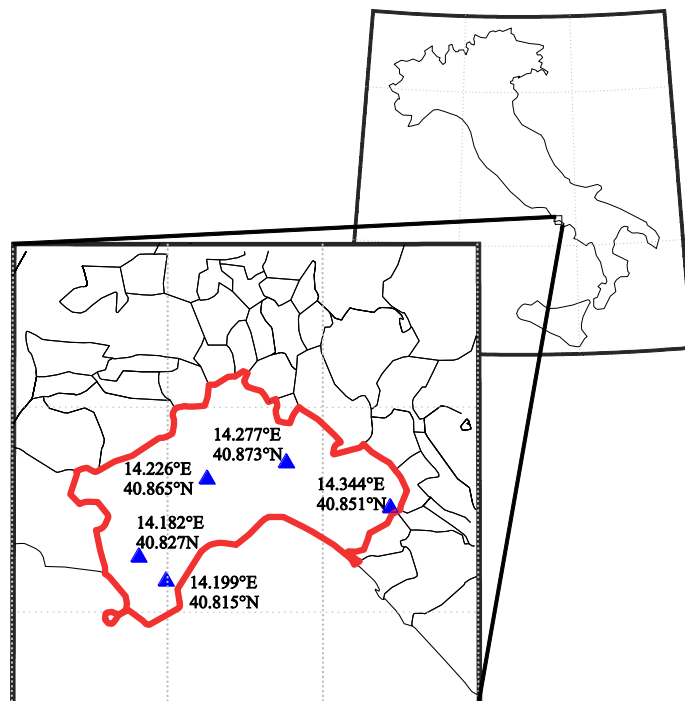
472 It is to also highlight that, although the vector collecting sites threshold in MSPSHA can be completely
473 defined by the user, REASSESS allows to define the threshold vector from the results of single-site PSHA.
474 For example, the thresholds can be chosen as the values with the same exceedance return period at each site
475 according to single-site PSHA, as illustrated in one of the examples below.

476 7. Illustrative examples

477 Some examples of the analyses REASSESS V2.0 enables are illustrated herein. To this aim five sites are
478 considered; incidentally, they correspond to the five main hospitals of the health infrastructure for municipality
479 of Naples (Italy): *Ospedale del Mare*, *San Giovanni Bosco*, *Cardarelli*, *San Paolo* and *Fatebenefratelli* (see
480 Figure 6 in which the sites and the municipality boundaries are highlighted). The inter-site distance ranges
481 between two and thirteen kilometers.

482 The following sections refer to the results of PSHA, SPSHA and MSPSHA. For all of them, the Meletti *et*
483 *al.* (2008) – *Magnitude rates from DPC-INGV-SI- Branch 921* source model is used (see Section 5). For the
484 aim of this paper, all the sites are assumed on rock soil conditions. In the case of SPSHA, the selected model
485 defining parameters of Equation (13) is Lolli and Gasperini (2003). All the data represented in the figures are

486 taken from the texts files automatically saved by REASSESS (to assemble the figures of the paper, the format
487 of the plots is slightly different from the one of the software).



488
489 Figure 6. Geographical location of the sites within the municipality of Naples.

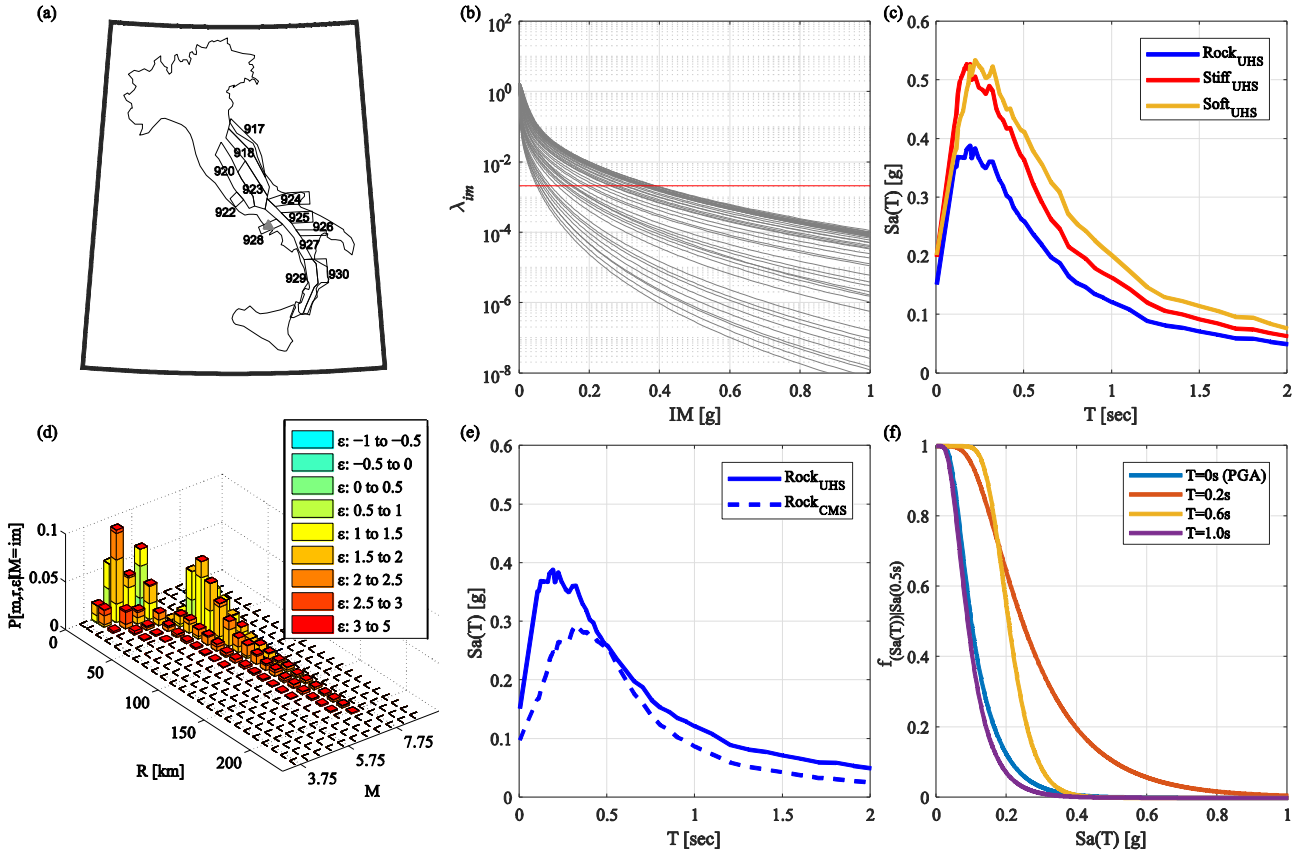
490 7.1.Single-site PSHA

491 Because the considered sites can be considered close from the seismic hazard assessment point of view,
492 differences in terms of single-site analysis, are minor. Thus, only one of the locations is considered for PSHA
493 and SPSHA: 14.277°E, 40.873°N. Figure 7 summarizes the result of single-site PSHA computed for the site.

494 In Figure 7a it is reported the location of the site (grey triangle) and the twelve seismic zones (out of the
495 thirty-six in total, numbered from 901 to 936) of the model of Meletti et al. (2008) contributing to the hazard
496 are plotted (these zones are automatically identified by REASSESS among those of the selected database).
497 Figure 7b reports the hazard curves computed for the whole forty-seven spectral periods of the GMPE. In the
498 same plot, the annual rate of exceedance equal to 0.0021, corresponding to the 475 return period (T_R) of
499 exceedance, is also plotted (red horizontal line). This return period is the one for which are computed the
500 UHS's in Figure 7c (the three soil conditions allowed by the GMPE are considered; i.e., *rock*, *stiff* and *soft*
501 soil). Such spectra have a peak ground acceleration (PGA) equal to about 0.2g and are representative of a
502 medium-high hazard site in Italy (see Stucchi et al., 2011).

503 Selecting as IM the pseudo-acceleration response spectral ordinate at 0.5s period, $Sa(0.5s)$, the occurrence
 504 disaggregation for a return period of 475 years is reported in Figure 7d. Such a disaggregation (for the
 505 occurrence of im) is computed as per Equation (6); however, because RVs are represented in a discretized way
 506 assuming bins of 10km distance and 0.5 magnitude, the PDF, $f_{M,R,\varepsilon|IM}$, is rendered in the plot by the
 507 corresponding discretized form, $P[m,r,\varepsilon|im]$. Disaggregation distribution is bi-modal, being the disaggregated
 508 hazard mainly affected by two seismic zones: the one in which the site is enclosed to (namely zone 928) and
 509 the zone 927 that, although is more distant than 928, is able to generate higher magnitude events and more
 510 frequently (see Iervolino et al., 2011, for a deeper discussion).

511 The CMS is reported in Figure 7e: conditioning IM is maintained $Sa(0.5s)$ corresponding to $T_R = 475$
 512 years. Finally, the conditional hazard distribution, Equation (7), for four pseudo-spectral accelerations at 0
 513 (PGA), 0.2s, 0.6s and 1.0s conditional to the same primary IM are reported in Figure 7f.



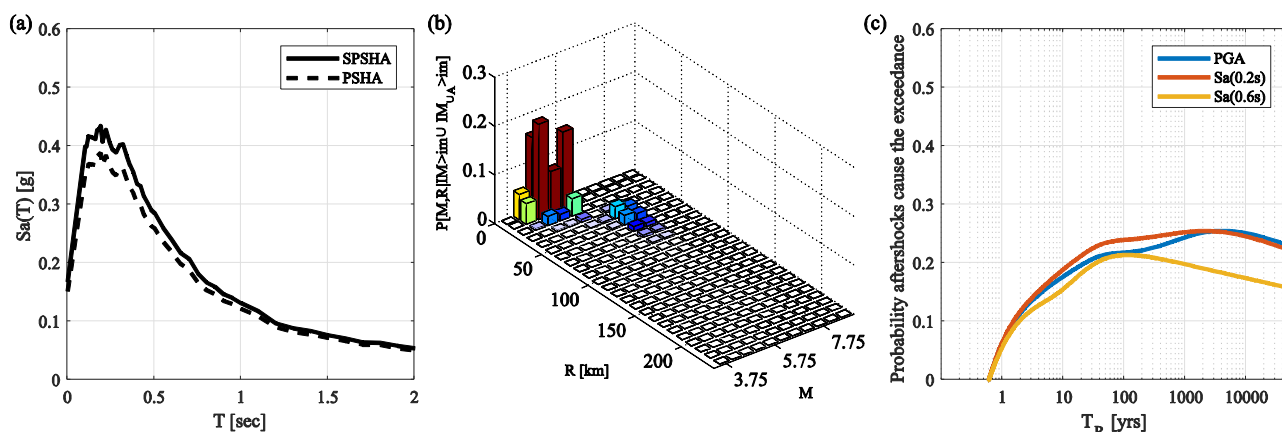
514 Figure 7. (a) Geographic location of the site and areal sources contributing to the hazard; (b) hazard curves (grey lines)
 515 computed for all the spectral period provided by the GMPE and annual rate corresponding to the 475 return period (red
 516 line); (c) UHS' with a 475 years return period; (d) hazard disaggregation distribution for the occurrence of the $Sa(0.5s)$
 517

518 with a 475 years return period; (e) CMS and (f) conditional hazard distributions assuming as primary IM the $Sa(0.5s)$
 519 with a 475 years return period.

520 7.2. SPSHA

521 For the same site as Section 7.1, Figure 8a shows the UHS' corresponding to 475 years return period and
 522 computed via both SPSHA and PSHA. The latter case corresponds to classical hazard, while the former
 523 includes the effect of aftershocks. Sequence's effect produces a maximum increase of UHS from PSHA equal
 524 to 12% which corresponds to a vibration period equal to 0.1s. However, increments over the whole range of
 525 analysed periods are equal or higher than 7%; the minimum value is recorded at 1.5s.

526 Both kinds of sequence-based disaggregation are also computed. The mainshock magnitude and distance
 527 disaggregation distribution, that is Equation (15), is shown for the PGA intensity measure and 475 years
 528 exceedance return period (Figure 8b); it is interesting to note that, accounting for the sequence modifies the
 529 proportion between first and second modal values with respect to Figure 7d (Chioccarelli et al., 2018).



530 Figure 8. (a) Comparison among UHS from PSHA and SPSHA for a 475yr return period; (b) Hazard disaggregation
 531 distribution of PGA with a 475 years return period; (c) aftershock disaggregation for PGA, $Sa(0.2s)$ and $Sa(0.6s)$.
 532 Figure 8c provides the aftershock disaggregation, Equation (16), performed for three IMs: PGA, $Sa(0.2s)$ and

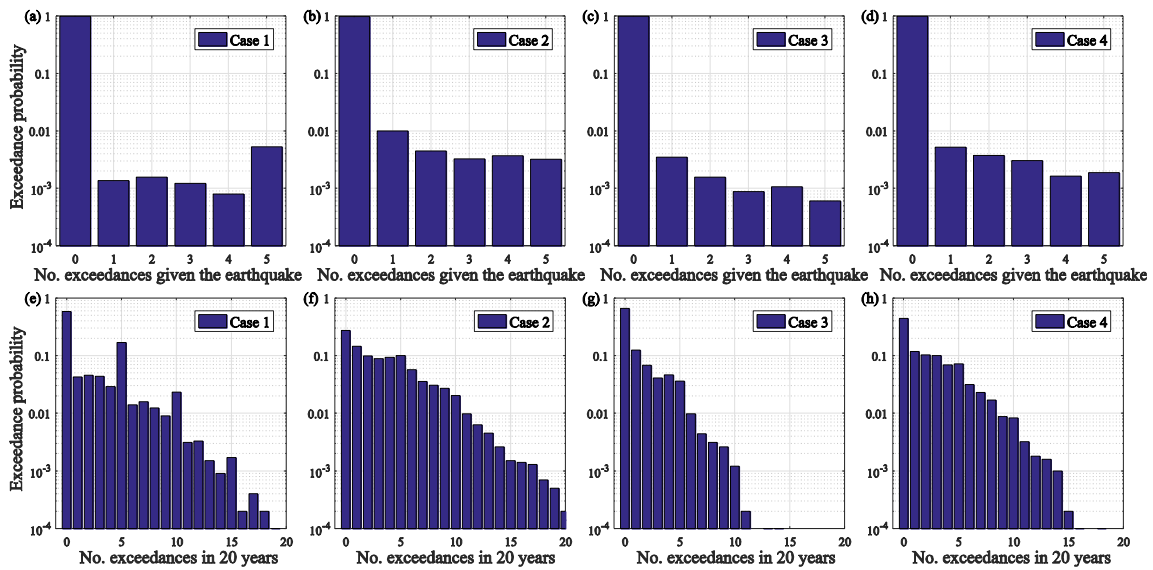
533 $Sa(0.6s)$. Aftershock disaggregation is here represented as a function of the increasing return period even if
 534 output text files provide them as function of both IM thresholds and return period. All these disaggregation
 535 distributions have a non-monotonic trend. In fact, they start from zero because it can be verified that when im
 536 approaches zero, results of Equation (3) and Equation(11) are equal, i.e., aftershock has no effect. The
 537 maximum value of disaggregation for PGA is 0.26 corresponding to a return period of about 4000 years;
 538 maximum disaggregation for $Sa(0.2s)$ is 0.26 and it occurs for a return period of about 2000 years; finally,
 539 maximum disaggregation for $Sa(0.6s)$ is 0.21 and correspond to T_R of about 100 years. The non-monotonic
 540

541 trend of the plots indicates that the aftershock contribution to the hazard has a variable significance with the
 542 hazard threshold. Moreover, the different return period to which each disaggregation reaches its maximum
 543 suggests that aftershock effect is also dependent on the considered spectral period.

544 7.3. MSPSHA

545 Results of MSPSHA are reported in this section referring to the whole set of the five sites introduced above
 546 (Section 7). A set of five IMs has been selected for each of the site: PGA, $Sa(0.2s)$, $Sa(0.5s)$, $Sa(0.6s)$,
 547 $Sa(1.0s)$. Profiting of the REASSESS functionalities discussed in Section 6.3, the vector of IMs collecting
 548 the threshold values for each site, which is required for MSPSHA, is chosen in a way that the corresponding
 549 T_R , from single-site PSHA, are the same among all the sites: the common return period is, arbitrarily, 50 years.

550 The distribution of the number of exceedances at the sites given the occurrence of the event and the
 551 distribution of the number of exceedance collectively observed at the sites in a time window of 20 years are
 552 the output here, chosen among those available (see Section 6.2). Both types of distribution are computed
 553 referring to four different cases: in (1) at each of the five sites, PGA is the considered IM; in (2) and (3) the
 554 considered IM at the sites is $Sa(0.5s)$ and $Sa(1.0s)$, respectively; finally, in (4) a different intensity measure
 555 is selected at each site: PGA at site one, $Sa(0.2s)$ at site two, $Sa(0.5s)$ at site three, $Sa(0.6s)$ at site four and
 556 $Sa(1.0s)$ at site five. The MPF of the total number of exceedances given the occurrence of an earthquake is
 557 reported in the first line of panels of Figure 9, from (a) to (d) corresponding to cases from 1 to 4, respectively.



559 Figure 9. MPF of the total number of exceedances at the sites (a) given the event and (b) in 20 years.
560 This result is representative of a specific case scenario which corresponds to the occurrence of a generic event.
561 It appears that the most probable number of exceedance is zero while the exceedance probabilities at one, more
562 than one, or all the sites are of the same order of magnitude.

563 The second line of the figure, i.e., plots from (e) to (h), shows the MPF of the total number of exceedances
564 in 20 years. It should be noted that the first five bars of the distributions show a trend similar to those observed
565 for the corresponding distributions conditional to the occurrence of an earthquake panels from (a) to (d).

566 8. Final remarks

567 REASSESS V2.0, a MATLAB-coded tool for probabilistic seismic hazard analysis, has been presented. It is
568 a standalone application which operates via GUI and/or template-based input files that has been developed to
569 enable classical and advanced probabilistic seismic hazard assessment procedures. It is oriented towards
570 earthquake engineering practice.

571 In the paper, the basics of probabilistic seismic hazard analyses embedded in the software have been
572 recalled first, then implemented algorithms and the workflow of REASSESS have been discussed. The
573 software allows the user to define the input of the analyses in terms of: site(s) coordinates, GMPEs (selected
574 among an embedded database), intensity measures of interest, seismic sources (user-defined three-dimensional
575 faults, seismic sources (areal) zones, or sources selected from embedded databases), and structure of logic tree,
576 if any.

577 When single-site analyses are of concern, REASSESS is able to provide classical results of PSHA such as
578 hazard curves, even in terms of spectral-shape-based (i.e., advanced) ground motion intensity measures.
579 Moreover, uniform hazard and conditional mean spectra, together with disaggregation distributions given the
580 occurrence or the exceedance of the IM threshold, can be computed. Conditional hazard can also be computed
581 for PGV or pseudo-spectral accelerations selected as secondary intensity measures. Moreover, single-site
582 analyses may also be performed accounting for the effect of the aftershocks. With this type of analysis, named
583 SPSHA, available output is represented by: hazard curves, UHS, magnitude-distance disaggregation
584 distribution and aftershock disaggregation. PSHA and SPSHA are implemented taking advantage of the
585 accuracy and low computational demand allowed by matrix calculus of MATLAB.

586 For portfolio of sites that can be subjected to the same seismic sources, the software is able to perform the
587 so-called MSPSHA providing, for a vector of IM thresholds, different probabilistic results all related to the
588 exceedances possibly observed the sites. A two-step simulation algorithm to carry out MSPSHA, allows to
589 profit of random field simulations of IMs in generic earthquakes.

590 REASSES was optimized for accuracy of numerical computation, analysis time and ease of use, which was
591 illustrated herein via a few applications, not exhaustive of the software capabilities. To this aim it also
592 implements calculation shortcuts and provides a series of options of input/output management. It is finally to
593 note that a practical user guide (tutorial) can be found online at
594 http://wpage.unina.it/iuniervo/doc_en/REASSESS.htm, which is the same site where the software is available
595 under a Creative Commons license: attribution—non-commercial—non derived.

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